The use of appropriate calibration curves corrects for systematic differences in liver R2* values measured using different software packages

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Keywords

liver iron concentration; Magnetic Resonance Imaging; R2*; decay model

Precise, reproducible measurements of liver iron concentration (LIC) are critical for the early diagnosis, treatment and follow-up of patients with primary or secondary iron overload. Magnetic resonance imaging (MRI), which exploits the paramagnetic properties of iron, has gained acceptance as a noninvasive and accurate tool for LIC assessment. Two different techniques have been described: methods measuring signal intensity ratio between liver and muscles (Gandon, *et al* 2004), and relaxometry methods measuring absolute R2 (St Pierre, *et al* 2005) or R2* (Wood, *et al* 2005) values, which increase proportionally to iron concentration. The gradient-echo R2* technique is most widely used in clinical practice (Anderson, *et al* 2001, Hankins, *et al* 2009, Meloni, *et al* 2011, Wood, *et al* 2005), because it is faster and easier than R2 acquisition.

R2* values can be converted to underlying tissue iron concentration, using appropriate calibration curves. The first calibration curve was proposed by Anderson *et al* (2001) and later updated by Garbowski *et al* (2009). This methodology has been implemented in a popular analysis software named "ThalassaemiaTools", a CMRtools (Cardiovascular

Imaging Solutions Ltd, London, UK) plug-in. A second calibration curve was proposed (Wood *et al* 2005) and later confirmed (Hankins *et al* 2009), but it demonstrated a 15% lower scaling coefficient between R2* and LIC.

The two R2* analysis methods differ in the size of analysed region of interest (ROI) and the model used to fit the signal decay at different echo times. A systematic comparison of R2* and LIC values obtained with the two methodologies has never been reported. This study aimed to detect the potential differences in R2* values obtained with different post-processing approaches and to explore whether the detected differences would be corrected when converted into LIC values.

Single- and multi-centre patient cohorts were used. The single-centre cohort included 45 patients (25 males, 16.4±10.2 years) scanned at the Children's Hospital of Los Angeles (CHLA). The multi-centre cohort (N=47; 19 females, 28.1±8.9 years) was baseline data from a phase II clinical trial of the iron chelator FBS0701 and was included to obtain higher generalizability

The study was approved by the CHLA Committee for the Protection of Human Subjects and the institutional review boards of all participating hospitals.

With the ThalassaemiaTools, a ROI was defined in an area of homogeneous liver tissue. All pixels were averaged together and fit to a single-exponential model. Later echo times were manually excluded from the fit in images where iron-mediated signal loss was high (Figures 1a–1b). R2* values (=1000/T2*) were converted into LIC by (Garbowski, *et al* 2009):

$$LIC_{Pennell} = 0.03 \cdot R2^*_{Pennell} + 0.7 \quad \text{[Equation 1]}$$

In our laboratory R2* measurements were performed using a custom-written software. The ROI included the entire liver profile in the slice, excluding the major hilar vessels. The signal in each pixel was fit to an exponential-plus-constant model, producing a R2* map (Figure 1c). The mean was calculated (R2*$_{Wood}$). LIC values (LIC$_{Wood}$) were calculated as:

$$LIC_{Wood} = 0.0254 \cdot R2^*_{Wood} + 0.2 \quad \text{[Equation 2]}$$

R2* assessment, by either method, is limited to LIC<40 mg/g because rapid signal decay precludes adequate characterization of the relaxation curve for higher iron concentrations.

In order to compare the two approaches linear regression analysis and Bland-Altman technique were used.

The results are indicated in Table I and Supplemental Figures 1, 2.

For both the cohorts the relationship between R2*$_{Wood}$ and R2*$_{Pennell}$ values was well described by a line. Results were unbiased for R2*<300 Hz, but large systematic differences in R2* appeared at higher values, with the exponential-plus-constant fits averaging ~20% higher. For the single-centre cohort the mean difference (± standard deviation) was 54.7±85.7 Hz (95% confidence intervals [CI], 28.9–80.5 Hz), corresponding to a percentage

R2* difference of 9.1±11.8%. Individually, 41 patients (93%) had $R2*_{Wood} > R2*_{Pennell}$ value. For the multi-centre cohort the mean difference was 59.5±76.7 Hz (95%CI, 36.9–81.9 Hz), corresponding to a relative difference of 8.5±13.8%.

When the technique-appropriate calibration curves were used, this bias effectively disappeared, producing excellent agreement between the two approaches. For the single-centre cohort the mean difference was −0.8±1.5 mg/g/dry weight (dw) (95%CI, −1.3 to −0.3 mg/g/dw). Individual LIC estimates had 95%CI from −3.8 to 2.2 mg/g/dw. For the multi-centre cohort the mean difference was −1.0±1.4 mg/g/dw (95%CI, −1.4 to −0.6 mg/g/dw), indicating a small, systematic bias. Individual LIC estimates had 95%CI of −3.8 to 1.8 mg/g/dw.

There is ongoing discussion as to which of the two decay models most closely describes the true tissue relaxation. By simulation (true value known), the key determinant is whether there is any other signal contribution besides iron. If there is any background signal from non-iron containing tissue, the exponential-plus-constant model is more accurate (Positano, *et al* 2009). If there is no signal contribution from bile, blood or fat, the truncated exponential is more appropriate (Beaumont, *et al* 2009).

Given that the two approaches sample the liver differently, one could potentially attribute the observed R2* differences to systematic gradients in iron distribution with proximity to the hilum. This was demonstrated to be untrue (McCarville, *et al* 2010, Positano, *et al* 2009).

The 0.8–1 mg/g residual bias between the two LIC estimates is clinically irrelevant at higher values, but could be important when assessing the risk of over chelation. Some of the bias is evident from inspection of Equations 1 and 2; the calibration curves have a 0.5 mg/g difference in y-intercept. To ensure good extrapolation into normal iron levels, the y-intercept in Equation 2 was constrained such that the calibration curve passes through the middle of the normal range (Wood, *et al* 2005), while Equation 1 was not. To illustrate the potential consequences of this difference, R2* values in healthy volunteers are typically reported to be 30–40 Hz. Equation 2 predicts LIC values of 1.0–1.2 mg/g and Equation 1 predicts LIC values of 1.4–1.6 mg/g. Practitioners who use Equation 1 should know that "normal" LIC corresponds to a value around 1.5 mg/g and base their therapeutic judgments about that set-point.

In conclusion, both signal decay models yield clinically-acceptable estimates of LIC if the ROI's are drawn correctly and the proper calibration curve is applied to correct for systematic differences in R2* estimation. Proper choice of technique at any given institution will depend on software availability and training. However, in the literature, R2* values should be converted into LIC values using the appropriate calibration curve to facilitate comparisons across studies.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

ACKNOWLEDGMENTS

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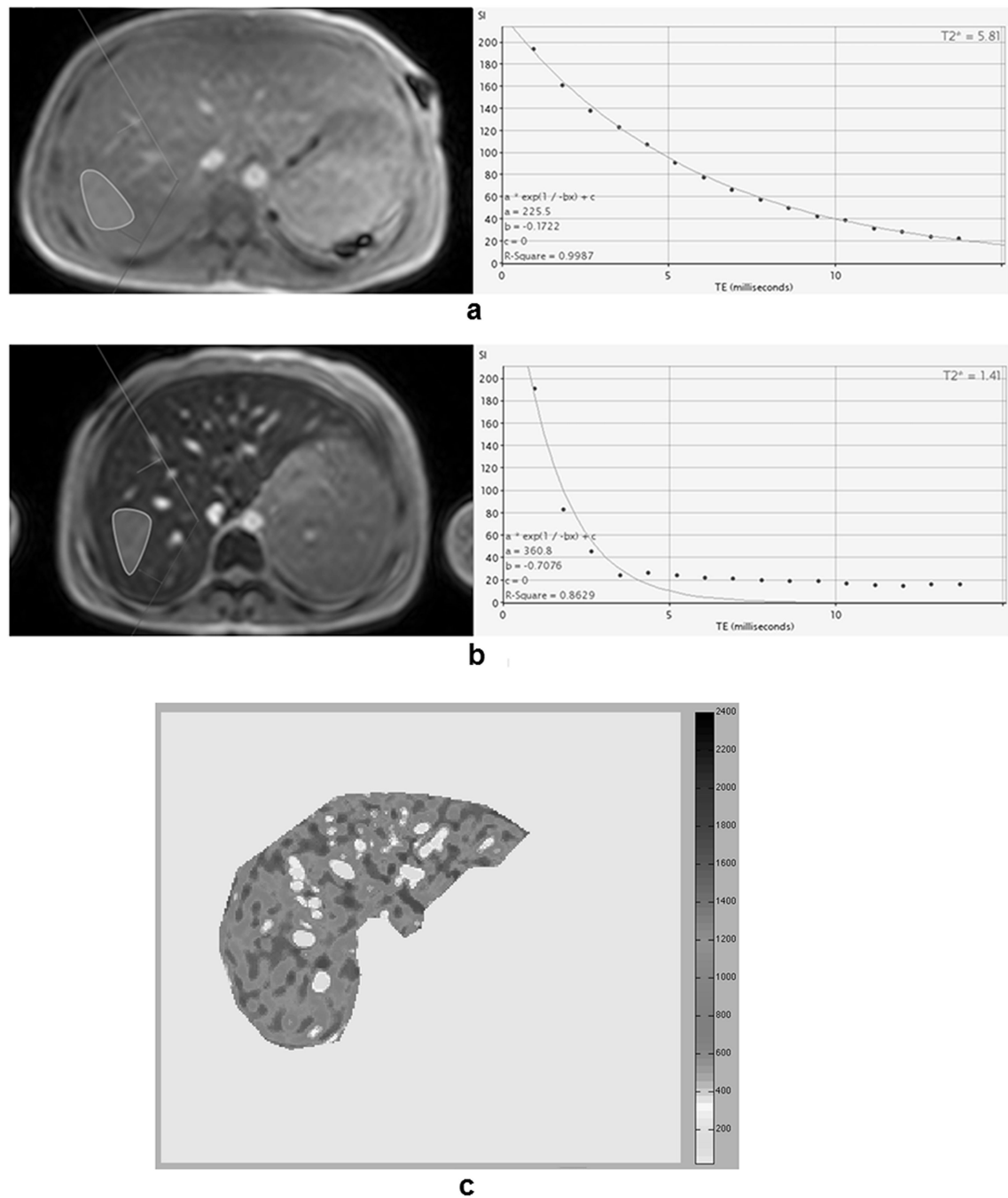
Figure 1.
a) and b) Screenshot of CMRtools software (region of interest definition in the left and signal decay fitting in the right) for a patient with moderate iron overload and for a patient with severe iron overload, respectively. For the patient with severe iron overload, the single-exponential model was not able to compensate the signal offset. In order to obtain an optimal fitting (R-square > 0.99), the last echo times (TEs) were excluded by de-selecting them. The final T2* was 1.05 ms, corresponding to a R2* of 952.4 Hz and a liver iron concentration (LIC) of 29.3 mg/g/dw. c) R2* map obtained with the Wood's approach for

the same patient in Figure 1b. The mean R2* for this distribution was 1087.9 Hz, corresponding to a LIC of 29.5 mg/g/dw.

Table 1

Comparison between $R2^*_{Wood}$ and $R2^*_{Pennell}$ values and between LIC_{Wood} and $LIC_{Pennell}$ values for the two cohorts of patients. LIC values are expressed also in µmol/g/dw, calculated from values expressed in mg/g/dw.

| | Values | Regression Analysis | | | Bland Altman | |
	Mean ± standard deviation [range]	Slope	Intercept	R-squared	Mean difference	Limits
Single-centre cohort						
$R2^*_{Wood}$ vs $R2^*_{Pennell}$	422.3±445.6 Hz [29.7–1344.9 Hz] vs 367.5±380.6 Hz [28.1–1219.5 Hz]	1.16±0.02	−3.99±12.72 Hz	0.982	54.7 Hz	−113.3 to 222.8 Hz
LIC_{Wood} vs $LIC_{Pennell}$	10.9±11.3 mg/g/dw [0.9–34.4 mg/g/dw] vs 11.7±11.4 mg/g/dw [1.5–37.3 mg/g/dw]	0.98±0.02	−0.59±0.33 mg/g/dw	0.991	−0.8 mg/g/dw	−3.8 to 2.2 mg/g/dw
LIC_{Wood} vs $LIC_{Pennell}$	195.6±202.6 µmol/g/dw [17.1–615.1 µmol/g/dw] vs 209.9±204 µmol/g/dw [27.6–667.4 µmol/g/dw]	0.98±0.02	−10.54±5.97 µmol /g/dw	0.991	−14.3 µmol /g/dw	−68.6 to 40.0 µmol /g/dw
Multi-centre cohort						
$R2^*_{Wood}$ vs $R2^*_{Pennell}$	499.8±314.2 Hz [64.2–1237.6 Hz] vs 440.4 ± 256.1Hz [62.2–1087.0 Hz]	1.21± 0.02	−32.07±16.31 Hz	0.969	59.5 Hz	−90.8 to 209.7 Hz
LIC_{Wood} vs $LIC_{Pennell}$	12.9±7.9 mg/g/dw [1.8–31.6 mg/g/dw] vs 13.9±7.7 mg/g/dw [2.5–33.3 mg/g/dw]	1.02±0.03	−1.33±0.43 mg/g/dw	0.969	−1.0 mg/g/dw	−3.8 to 1.8 mg/g/dw
LIC_{Wood} vs $LIC_{Pennell}$	230.9±142.9 µmol/g/dw [32.8–566.3 µmol/g/dw] vs 249.0±137.5 µmol/g/dw [45.9–596.2 µmol/g/dw]	1.02±0.03	−23.82±7.71 µmol /g/dw	0.969	−18.2 µmol /g/dw	−67.7 to 31.4 µmol /g/dw